

Example 3: Time-Invariant Predictors of Practice Effects (uses same data as Example 2)

In this example we will examine time-invariant predictors of individual differences in intercepts, linear slopes, and quadratic slopes representing improvement in RT (in msec) across six practice sessions. We will examine age, abstract reasoning, and education in sequential conditional (predictor) models.

SAS Code for Data Manipulation:

```
* Centering level-2 predictor variables for analysis;
DATA work.example23; SET work.example23;
  age80 = age - 80;          * Convenient value;
  reas22 = absreas - 22;    * Near sample mean;
  LABEL age80 = "age80: Age Centered (0=80)"
         reas22 = "reas22: Abstract Reasoning Centered (0=22)";
  * Make education a grouping variable for purpose of demonstration only;
  IF educyrs LE 12          THEN educgrp=1;
  ELSE IF educyrs GT 12 AND EducYrs LE 16 THEN educgrp=2;
  ELSE IF educyrs GT 16    THEN educgrp=3;
  ELSE IF educyrs = .      THEN educGrp=.;
  LABEL educgrp = "educgrp: Education Group (1=HS, 2=BA, 3=GRAD)";

* Removing cases with missing predictors;
  IF NMISS(age80, reas22, educgrp)>0 THEN DELETE;
RUN;
```

SPSS Code for Data Manipulation:

```
* Centering level-2 predictor variables for analysis.
DATASET ACTIVATE example23 WINDOW=FRONT.
COMPUTE age80 = age - 80.
COMPUTE reas22 = absreas - 22.
VARIABLE LABELS
  age80 "age80: Age Centered (0=80)"
  reas22 "reas22: Abstract Reasoning Centered (0=22)".
* Make education a grouping variable for purpose of demonstration only.
IF educyrs LE 12          educgrp=1.
IF educyrs GT 12 AND educyrs LE 16 educgrp=2.
IF educyrs GT 16        educgrp=3.
VARIABLE LABELS educgrp "educgrp: Education Group (1=HS, 2=BA, 3=GRAD)".

* Removing cases with missing predictors.
SELECT IF (NVALID(age80, reas22, educgrp)=3).
EXECUTE.
```

STATA Code for Data Manipulation:

```
* centering level-2 predictor variables for analysis
gen age80 = age - 80
gen reas22 = absreas - 22
label variable age80 "age80: Age Centered (0=80 years)"
label variable reas22 "reas22: Abstract Reasoning Centered (0=22)"
* make education a grouping variable for purpose of demonstration only
gen educgrp=.
replace educgrp=1 if (educyrs <= 12)
replace educgrp=2 if (educyrs > 12 & educyrs <= 16)
replace educgrp=3 if (educyrs > 16)
label variable educgrp "educgrp: Education Group (1=HS, 2=BA, 3=GRAD)"

* create new variable to hold number of missing cases
* then drop cases with incomplete predictors
egen nummiss = rowmiss(age80 reas22 educgrp)
drop if nummiss>0
```

Model 3b. Random Quadratic Time Baseline (in ML now)

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$$

Level 2:

$$\text{Intercept: } \beta_{0i} = \gamma_{00} + U_{0i}$$

$$\text{Linear: } \beta_{1i} = \gamma_{10} + U_{1i}$$

$$\text{Quadratic: } \beta_{2i} = \gamma_{20} + U_{2i}$$

```
TITLE1 "SAS Model 3b: Random Quadratic Time Baseline in ML";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 IC METHOD=ML;
  CLASS ID session;
  MODEL nm3rt = c1sess c1sess*c1sess / SOLUTION DDFM=Satterthwaite OUTPM=work.TimePred;
  RANDOM INTERCEPT c1sess c1sess*c1sess / G GCORR V VCORR TYPE=UN SUBJECT=ID;
  REPEATED session / R TYPE=VC SUBJECT=ID; RUN;
PROC CORR NOSIMPLE DATA=work.TimePred; VAR nm3rt pred; RUN;
```

```
TITLE "SPSS Model 3b: Random Quadratic Time Baseline in ML".
MIXED nm3rt BY ID session WITH c1sess
  /METHOD = ML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess c1sess*c1sess
  /RANDOM = INTERCEPT c1sess c1sess*c1sess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /SAVE = FIXPRED (predtime).
CORRELATIONS nm3rt predtime.
```

```
* STATA Model 3b: Random Quadratic Time Baseline in ML
xtmixed nm3rt c.c1sess c.c1sess#c.c1sess, || id: c1sess c1sess2, ///
  variance ml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),
  estat recovariance, level(id),
  estimates store Baseline, // save LL for LRT
  predict predtime // save fixed-effect predicted outcomes
corr nm3rt predtime // get total r to make r2
```

STATA output:

```
Mixed-effects ML regression          Number of obs   =       606
Group variable: id                  Number of groups =       101
                                     Obs per group:  min =         6
                                     avg   =       6.0
                                     max   =         6
                                     Wald chi2(2)     =       72.45
Log likelihood = -4160.8833          Prob > chi2     =       0.0000
```

nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
c1sess	-120.8999	19.94803	-6.06	0.000	-159.9973	-81.80251
c.c1sess#c.c1sess	13.86561	3.398459	4.08	0.000	7.204756	20.52647
_cons	1945.85	53.58259	36.31	0.000	1840.83	2050.87

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]	
id: Unstructured				
var(c1sess)	25437.86	5781.419	16293.81	39713.52
var(c1sess2)	622.8	169.99	364.7687	1063.358
var(_cons)	273306.9	40831.76	203930.4	366285.1
cov(c1sess,c1sess2)	-3837.723	968.8047	-5736.545	-1938.9
cov(c1sess,_cons)	-35261.67	11771.5	-58333.38	-12189.95
cov(c1sess2,_cons)	3845.378	1921.468	79.37031	7611.386

```
var(Residual) |      20298.2   1649.119       17310.2   23801.98
```

```
-----
LR test vs. linear regression:      chi2(6) =   891.99   Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.
```

```
.      estat ic, n(101),
```

```
-----
Model |   Obs   ll(null)   ll(model)   df       AIC       BIC
-----+-----
. |   101           . -4160.883   10   8341.767   8367.918
-----
```

In ML, the #parms is ALL parms (both sides of model). So STATA's versions should agree with other programs.

```
Note: N=101 used in calculating BIC
```

```
-----
      |   nm3rt   preptime
-----+-----
nm3rt |   1.0000
pretime | 0.1917   1.0000
```

R = .1917, so R² for time = .0367

The model for the means (fixed linear and quadratic session effects so far) accounted for ~4% of the variance in RT.

Model 4a. Age as Predictor of Intercept, Linear, and Quadratic Time Slopes

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$$

Level 2:

$$\text{Intercept: } \beta_{0i} = \gamma_{00} + \gamma_{01}(\text{Age}_i - 80) + U_{0i}$$

$$\text{Linear: } \beta_{1i} = \gamma_{10} + \gamma_{11}(\text{Age}_i - 80) + U_{1i}$$

$$\text{Quadratic: } \beta_{2i} = \gamma_{20} + \gamma_{21}(\text{Age}_i - 80) + U_{2i}$$

```
TITLE1 "SAS Model 4a: Age as Predictor of Intercept, Linear, and Quadratic";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 IC METHOD=ML;
CLASS ID session;
MODEL nm3rt = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
              / SOLUTION DDFM=Satterthwaite OUTPM=work.AgePred;
RANDOM INTERCEPT c1sess c1sess*c1sess / G GCORR TYPE=UN SUBJECT=ID;
REPEATED session / TYPE=VC SUBJECT=ID;
* Requesting additional effects for age;
ESTIMATE "Age Effect at Session 1" age80 1 c1sess*age80 0 c1sess*c1sess*age80 0;
ESTIMATE "Age Effect at Session 2" age80 1 c1sess*age80 1 c1sess*c1sess*age80 1;
ESTIMATE "Age Effect at Session 3" age80 1 c1sess*age80 2 c1sess*c1sess*age80 4;
ESTIMATE "Age Effect at Session 4" age80 1 c1sess*age80 3 c1sess*c1sess*age80 9;
ESTIMATE "Age Effect at Session 5" age80 1 c1sess*age80 4 c1sess*c1sess*age80 16;
ESTIMATE "Age Effect at Session 6" age80 1 c1sess*age80 5 c1sess*c1sess*age80 25;
RUN; PROC CORR NOSIMPLE DATA=work.AgePred; VAR nm3rt pred; RUN;
```

```
TITLE "SPSS Model 4a: Age as Predictor of Intercept, Linear, and Quadratic".
MIXED nm3rt BY ID session WITH c1sess age80
  /METHOD = ML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
  /RANDOM = INTERCEPT c1sess c1sess*c1sess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /SAVE = FIXPRED (predage)
  /TEST = "Age Effect at Session 1" age80 1 c1sess*age80 0 c1sess*c1sess*age80 0
  /TEST = "Age Effect at Session 2" age80 1 c1sess*age80 1 c1sess*c1sess*age80 1
  /TEST = "Age Effect at Session 3" age80 1 c1sess*age80 2 c1sess*c1sess*age80 4
  /TEST = "Age Effect at Session 4" age80 1 c1sess*age80 3 c1sess*c1sess*age80 9
  /TEST = "Age Effect at Session 5" age80 1 c1sess*age80 4 c1sess*c1sess*age80 16
  /TEST = "Age Effect at Session 6" age80 1 c1sess*age80 5 c1sess*c1sess*age80 25.
CORRELATIONS nm3rt predage.
```

```
* STATA Model 4a: Age as Predictor of Intercept, Linear, and Quadratic
xtmixed nm3rt c.c1sess c.c1sess#c.c1sess          ///
      c.age80 c.age80#c.c1sess c.age80#c.c1sess#c.c1sess,  ///
      || id: c1sess c1sess2,                          ///
      variance ml covariance(un) residuals(independent,t(session)),
      estat ic, n(101),
      estat recovariance, level(id),
      estimates store age,          // save LL for LRT
      lrtest Age Baseline,         // LRT against non-age baseline
      predict predage              // save fixed-effect predicted outcomes
      margins, at(c.c1sess=(0(1)5)) dydx(c.age80) vsquish // age slope per session
      margins, at(c.c1sess=(0(1)5) c.age80=(-5 0 5)) vsquish // predictions per session
      marginsplot, name(predicted_age, replace) // plot age predictions
corr nm3rt predage // get total r to make r2
```

STATA output:

```
Mixed-effects ML regression          Number of obs   =      606
Group variable: id                  Number of groups =      101
                                     Obs per group: min =       6
                                     avg =          6.0
                                     max =           6
                                     Wald chi2(5)      =     88.55
Log likelihood = -4155.1009          Prob > chi2     =     0.0000
```

	nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	c1sess	-121.8325	19.66948	-6.19	0.000	-160.3839	-83.28099
	c.c1sess#c.c1sess	13.97744	3.375686	4.14	0.000	7.361221	20.59367
	age80	29.04954	8.377364	3.47	0.001	12.63021	45.46887
	c.age80#c.c1sess	-5.594634	3.251901	-1.72	0.085	-11.96824	.7789759
	c.age80#c.c1sess#c.c1sess	.6709122	.558093	1.20	0.229	-.42293	1.764754
	_cons	1950.692	50.67139	38.50	0.000	1851.378	2050.006

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]	
id: Unstructured				
var(c1sess)	24293.61	5623.947	15432.62	38242.33 → linear var down by 4.50%
var(c1sess2)	606.3449	167.7546	352.5508	1042.84 → quad var down by 2.64%
var(_cons)	242456.1	36492.45	180516.8	325648.3 → intercept var down 11.29%
cov(c1sess,c1sess2)	-3700.505	949.404	-5561.302	-1839.707
cov(c1sess,_cons)	-29320.18	10868.45	-50621.95	-8018.411
cov(c1sess2,_cons)	3132.873	1793.883	-383.0738	6648.819
var(Residual)	20298.2	1649.119	17310.2	23801.98 → residual var not reduced

```
LR test vs. linear regression:      chi2(6) = 857.76  Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.
```

```
.      estat ic, n(101),
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	101	.	-4155.101	13	8336.202	8370.198

Note: N=101 used in calculating BIC

```
.      estimates store Age,          // save LL for LRT
.      lrtest Age Baseline,         // LRT against non-age baseline
```

```
Likelihood-ratio test          LR chi2(3) = 11.56
(Assumption: Baseline nested in Age)  Prob > chi2 = 0.0090
```

Is the age model (4a) better than the baseline random quadratic model (3b)?
 Yes, $-2\Delta LL=11.6$ on $df=3$, $p=.009$

```
. predict predage // save fixed-effect predicted outcomes
(option xb assumed)
. margins, at(c1sess=(0(1)5)) dydx(age80) vsquish // age slope per session
```

Average marginal effects Number of obs = 606

```
Expression : Linear prediction, fixed portion, predict()
dy/dx w.r.t. : age80
1._at : c1sess = 0
2._at : c1sess = 1
3._at : c1sess = 2
4._at : c1sess = 3
5._at : c1sess = 4
6._at : c1sess = 5
```

These are the simple slopes for age at each session.

		Delta-method				
		dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]
age80	_at					
	1	29.04954	8.377364	3.47	0.001	12.63021 45.46887
	2	24.12582	7.609705	3.17	0.002	9.211068 39.04056
	3	20.54392	7.459286	2.75	0.006	5.923987 35.16385
	4	18.30385	7.330177	2.50	0.013	3.936962 32.67073
	5	17.4056	7.071475	2.46	0.014	3.545761 31.26543
	6	17.84917	7.054461	2.53	0.011	4.022683 31.67566

```
. margins, at(c1sess=(0(1)5) age80=(-5 0 5)) vsquish // predictions per session
```

Adjusted predictions Number of obs = 606

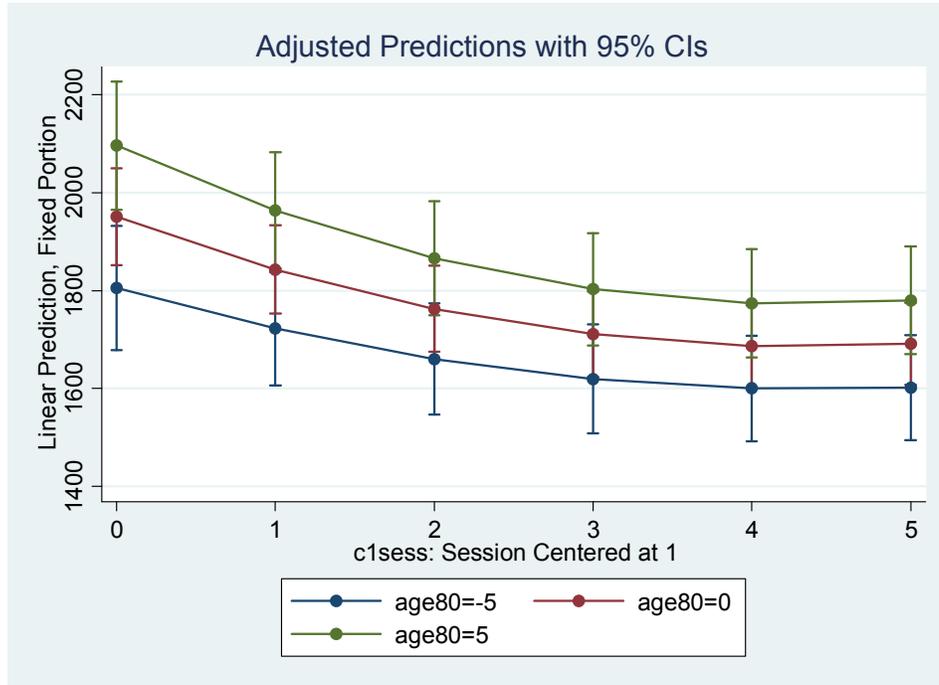
```
Expression : Linear prediction, fixed portion, predict()
1._at : c1sess = 0
       age80 = -5
2._at : c1sess = 0
       age80 = 0
3._at : c1sess = 0
       age80 = 5
```

(output continues for all other sessions)

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
_at						
	1	1805.444	64.84687	27.84	0.000	1678.347 1932.542
	2	1950.692	50.67139	38.50	0.000	1851.378 2050.006
	3	2095.94	66.62638	31.46	0.000	1965.354 2226.525
	4	1722.208	58.90463	29.24	0.000	1606.757 1837.659
	5	1842.837	46.02812	40.04	0.000	1752.623 1933.05
	6	1963.466	60.52108	32.44	0.000	1844.847 2082.085
	7	1660.217	57.74028	28.75	0.000	1547.048 1773.386
	8	1762.937	45.1183	39.07	0.000	1674.506 1851.367
	9	1865.656	59.32478	31.45	0.000	1749.382 1981.931
	10	1619.472	56.74089	28.54	0.000	1508.262 1730.682
	11	1710.991	44.33737	38.59	0.000	1624.092 1797.891
	12	1802.511	58.29796	30.92	0.000	1688.249 1916.773
	13	1599.973	54.73834	29.23	0.000	1492.688 1707.258
	14	1687.001	42.77258	39.44	0.000	1603.168 1770.834
	15	1774.029	56.24045	31.54	0.000	1663.8 1884.258
	16	1601.72	54.60664	29.33	0.000	1494.693 1708.747
	17	1690.966	42.66967	39.63	0.000	1607.335 1774.597
	18	1780.212	56.10514	31.73	0.000	1670.247 1890.176

The pattern of the interaction is shown by the simple effects of age at each session, graphed below.

```
. marginsplot, name(predicted_age, replace) // plot age predictions
```



Variables that uniquely identify margins: c1sess age80

```
. corr nm3rt predage // get total r to make r2
(obs=606)
```

	nm3rt	predage
nm3rt	1.0000	
predage	0.3269	1.0000

R = .3269, so R² for time+age = .1069
 The fixed effects of time before accounted for ~3.7% of the variance in RT, so there is a net increase of ~7% due to age.

Model 5a. +Abstract Reasoning as Predictor of Intercept, Linear, and Quadratic Time Slopes

Level 1: $y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$

Level 2:

Intercept: $\beta_{0i} = \gamma_{00} + \gamma_{01}(\text{Age}_i - 80) + \gamma_{02}(\text{Reason}_i - 22) + U_{0i}$

Linear: $\beta_{1i} = \gamma_{10} + \gamma_{11}(\text{Age}_i - 80) + \gamma_{12}(\text{Reason}_i - 22) + U_{1i}$

Quadratic: $\beta_{2i} = \gamma_{20} + \gamma_{21}(\text{Age}_i - 80) + \gamma_{22}(\text{Reason}_i - 22) + U_{2i}$

```
TITLE1 "SAS Model 5a: +Reasoning as Predictor of Intercept, Linear, and Quadratic";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 IC METHOD=ML;
  CLASS ID session;
  MODEL nm3rt = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
    reas22 c1sess*reas22 c1sess*c1sess*reas22
    / SOLUTION DDFM=Satterthwaite OUTPM=work.ReasPred;
  RANDOM INTERCEPT c1sess c1sess*c1sess / G GCORR TYPE=UN SUBJECT=ID;
  REPEATED session / TYPE=VC SUBJECT=ID;
RUN; PROC CORR NOSIMPLE DATA=work.ReasPred; VAR nm3rt pred; RUN;
```

```
TITLE "SPSS Model 5a: +Reasoning as Predictor of Intercept, Linear, and Quadratic".
MIXED nm3rt BY ID session WITH c1sess age80 reas22
  /METHOD = ML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
           reas22 c1sess*reas22 c1sess*c1sess*reas22
  /RANDOM = INTERCEPT c1sess c1sess*c1sess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /SAVE = FIXPRED (predreas).
CORRELATIONS nm3rt predreas.
```

```
* STATA Model 5a: +Reasoning as Predictor of Intercept, Linear, and Quadratic
xtmixed nm3rt c.c1sess c.c1sess#c.c1sess ///
  c.age80 c.age80#c.c1sess c.age80#c.c1sess#c.c1sess ///
  c.reas22 c.reas22#c.c1sess c.reas22#c.c1sess#c.c1sess, ///
  || id: c1sess c1sess2, ///
  variance ml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),
  estat recovariance, level(id),
  estimates store Reas, // save LL for LRT
  lrtest Reas Age, // LRT against age baseline
  predict predreas // save fixed-effect predicted outcomes
corr nm3rt predreas // get total r to make r2
```

STATA output:

```
Mixed-effects ML regression      Number of obs      =      606
Group variable: id              Number of groups   =      101
                                Obs per group: min =       6
                                avg =      6.0
                                max =       6
                                Wald chi2(8)      =     103.88
                                Prob > chi2      =      0.0000
Log likelihood = -4148.8645
```

	nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	c1sess	-119.7417	19.77414	-6.06	0.000	-158.4983	-80.98505
	c.c1sess#c.c1sess	13.30362	3.36557	3.95	0.000	6.707229	19.90002
	age80	22.27817	8.601751	2.59	0.010	5.419047	39.13729
	c.age80#c.c1sess	-6.492074	3.424732	-1.90	0.058	-13.20443	.2202772
	c.age80#c.c1sess#c.c1sess	.9601368	.5828914	1.65	0.100	-.1823093	2.102583
	reas22	-27.10041	11.11411	-2.44	0.015	-48.88366	-5.317155
	c.reas22#c.c1sess	-3.591742	4.425011	-0.81	0.417	-12.2646	5.081121
	c.reas22#c.c1sess#c.c1sess	1.157537	.7531395	1.54	0.124	-.3185897	2.633663
	_cons	1966.467	49.66585	39.59	0.000	1869.124	2063.811

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]			
id: Unstructured						
var(c1sess)	24040.63	5589.24	15242.24	37917.78	→ linear var down by 1.04%	
var(c1sess2)	580.0652	164.1907	333.0729	1010.216	→ quad var down by 4.33%	
var(_cons)	228049.3	34467.25	169581.6	306675.4	→ intercept var down by 5.94%	
cov(c1sess,c1sess2)	-3618.966	937.0759	-5455.601	-1782.331		
cov(c1sess,_cons)	-31229.47	10655.91	-52114.67	-10344.27		
cov(c1sess2,_cons)	3748.206	1747.738	322.7024	7173.709		
var(Residual)	20298.18	1649.113	17310.18	23801.94	→ residual var not reduced	

LR test vs. linear regression: chi2(6) = 832.43 Prob > chi2 = 0.0000

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	101	.	-4148.864	16	8329.729	8371.571

```

.      lrtest Reas Age,          // LRT against age baseline
Likelihood-ratio test          LR chi2(3) =    12.47
(Assumption: Age nested in Reas) Prob > chi2 =    0.0059

. corr nm3rt predreas          // get total r to make r2
(obs=606)
-----+-----
      nm3rt | 1.0000
predreas | 0.4011  1.0000

```

Is the reasoning model (5a) better than the age model (4a)?

Yes, $-2\Delta LL = 12.5$ on $df=3$, $p = .0059$, so ΔR^2 is significant

$R = .4011$, so R^2 for time+age+reas = .1609

The fixed effects of time and age before accounted for ~10.7% of the variance in RT, so there is a net increase of ~5.4% due to reasoning.

Model 5b. Abstract Reasoning on Intercept and Linear Time Slope Only

$$\text{Level 1: } y_{ti} = \beta_{0i} + \beta_{1i}(\text{Session}_{ti} - 1) + \beta_{2i}(\text{Session}_{ti} - 1)^2 + e_{ti}$$

Level 2:

$$\text{Intercept: } \beta_{0i} = \gamma_{00} + \gamma_{01}(\text{Age}_i - 80) + \gamma_{02}(\text{Reason}_i - 22) + U_{0i}$$

$$\text{Linear: } \beta_{1i} = \gamma_{10} + \gamma_{11}(\text{Age}_i - 80) + \gamma_{12}(\text{Reason}_i - 22) + U_{1i}$$

$$\text{Quadratic: } \beta_{2i} = \gamma_{20} + \gamma_{21}(\text{Age}_i - 80) + U_{2i}$$

```

TITLE1 "SAS Model 5b: Reasoning on Intercept and Linear Time Slope Only";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 IC METHOD=ML;
  CLASS ID session;
  MODEL nm3rt = cclass cclass*cclass age80 cclass*age80 cclass*cclass*age80
              reas22 cclass*reas22
              / SOLUTION DDFM=Satterthwaite OUTPM=work.ReasPred2;
  RANDOM INTERCEPT cclass cclass*cclass / G GCORR TYPE=UN SUBJECT=ID;
  REPEATED session / TYPE=VC SUBJECT=ID;
  * Requesting additional effects for reasoning instead;
  ESTIMATE "Reasoning Effect at Session 1" reas22 1 cclass*reas22 0;
  ESTIMATE "Reasoning Effect at Session 2" reas22 1 cclass*reas22 1;
  ESTIMATE "Reasoning Effect at Session 3" reas22 1 cclass*reas22 2;
  ESTIMATE "Reasoning Effect at Session 4" reas22 1 cclass*reas22 3;
  ESTIMATE "Reasoning Effect at Session 5" reas22 1 cclass*reas22 4;
  ESTIMATE "Reasoning Effect at Session 6" reas22 1 cclass*reas22 5;
RUN; PROC CORR NOSIMPLE DATA=work.ReasPred2; VAR nm3rt pred; RUN;

```

```

TITLE "SPSS Model 5b: +Reasoning as Predictor of Intercept, Linear Time Slope Only".
MIXED nm3rt BY ID session WITH cclass age80 reas22
  /METHOD = ML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = cclass cclass*cclass age80 cclass*age80 cclass*cclass*age80
          reas22 cclass*reas22
  /RANDOM = INTERCEPT cclass cclass*cclass | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /SAVE = FIXPRED (predreas2)
  /TEST = "Reasoning Effect at Session 1" reas22 1 cclass*reas22 0
  /TEST = "Reasoning Effect at Session 2" reas22 1 cclass*reas22 1
  /TEST = "Reasoning Effect at Session 3" reas22 1 cclass*reas22 2
  /TEST = "Reasoning Effect at Session 4" reas22 1 cclass*reas22 3
  /TEST = "Reasoning Effect at Session 5" reas22 1 cclass*reas22 4
  /TEST = "Reasoning Effect at Session 6" reas22 1 cclass*reas22 5.
CORRELATIONS nm3rt predreas2.

```

```

* STATA Model 5b: +Reasoning as Predictor of Intercept, Linear Time Slope Only
xtmixed nm3rt c.cclass c.cclass#c.cclass ///
  c.age80 c.age80#c.cclass c.age80#c.cclass#c.cclass ///
  c.reas22 c.reas22#c.cclass, || id: cclass cclass2, ///
  variance ml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),

```

```

estat recovariance, level(id),
estimates store Reas2, // save LL for LRT
lrtest Reas2 Age, // LRT against age baseline
margins, at(c.c1sess=(0(1)5)) dydx(c.reas22) vsquish // reas slope per session
margins, at(c.c1sess=(0(1)5) c.reas22=(-5 0 5)) vsquish // predictions per session
marginsplot, name(predicted_reas, replace) // plot reas predictions
predict predreas2 // save fixed-effect predicted outcomes
corr nm3rt predreas2 // get total r to make r2

```

STATA output:

```

Mixed-effects ML regression      Number of obs   =      606
Group variable: id              Number of groups =      101
                                Obs per group:  min =       6
                                avg =      6.0
                                max =       6
                                Wald chi2(7)      =     101.09
Log likelihood = -4150.032      Prob > chi2     =     0.0000

```

	nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	c1sess	-123.5416	19.82897	-6.23	0.000	-162.4057	-84.67758
	c.c1sess#c.c1sess	13.97744	3.375697	4.14	0.000	7.3612	20.59369
	age80	20.84705	8.561395	2.44	0.015	4.067021	37.62707
	c.age80#c.c1sess	-4.860993	3.290742	-1.48	0.140	-11.31073	1.588743
	c.age80#c.c1sess#c.c1sess	.6709122	.5580948	1.20	0.229	-.4229335	1.764758
	reas22	-32.82806	10.47071	-3.14	0.002	-53.35027	-12.30585
	c.reas22#c.c1sess	2.93618	1.241355	2.37	0.018	.5031693	5.369191
	_cons	1969.802	49.6827	39.65	0.000	1872.425	2067.178

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]			
id: Unstructured						
var(c1sess)	24876.74	5713.188	15860.12	39019.37	→ linear var up by -2.40%	
var(c1sess2)	606.3538	167.7562	352.557	1042.852	→ quad var not reduced	
var(_cons)	228693.2	34638.71	169952.4	307736.8	→ intercept var down by 5.68%	
cov(c1sess,c1sess2)	-3767.223	957.7847	-5644.446	-1889.999		
cov(c1sess,_cons)	-31963.12	10883.66	-53294.71	-10631.53		
cov(c1sess2,_cons)	3878.292	1787.961	373.9525	7382.631		
var(Residual)	20298.14	1649.108	17310.16	23801.9	→ residual var not reduced	

LR test vs. linear regression: chi2(6) = 830.58 Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.

```

.        estat ic, n(101),
-----+-----
Model | Obs    ll(null)    ll(model)    df        AIC        BIC
-----+-----
.     |    101            .    -4150.032    15       8330.064    8369.291

```

Note: N=101 used in calculating BIC

```

.        estimates store Reas2,        // save LL for LRT
.        lrtest Reas2 Age,            // LRT against age baseline

```

```

Likelihood-ratio test                LR chi2(2) =     10.14
(Assumption: Age nested in Reas2)    Prob > chi2 =     0.0063

```

```

.        margins, at(c1sess=(0(1)5)) dydx(reas22) vsquish    // reas slope per session

```

```

Average marginal effects                Number of obs    =        606

```

Expression : Linear prediction, fixed portion, predict()

Is the revised reasoning model (5b) still better than the age model (4a)?

Yes, $-2\Delta LL = 10.1$ on $df=2$, $p=.006$ (so only 2.4 of the previous $-2\Delta LL$ was due to reason*quad)

```
dy/dx w.r.t. : reas22
1._at       : c1sess      =      0
2._at       : c1sess      =      1
3._at       : c1sess      =      2
4._at       : c1sess      =      3
5._at       : c1sess      =      4
6._at       : c1sess      =      5
```

These are the simple slopes for reasoning at each session.

```
-----+-----
          |          Delta-method
          |          dy/dx  Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
reas22  _at |
  1 | -32.82806   10.47071   -3.14   0.002   -53.35027   -12.30585
  2 | -29.89188   9.961528   -3.00   0.003   -49.41612   -10.36764
  3 | -26.9557    9.586986   -2.81   0.005   -45.74585   -8.165552
  4 | -24.01952   9.363252   -2.57   0.010   -42.37116   -5.667884
  5 | -21.08334   9.301214   -2.27   0.023   -39.31338   -2.853295
  6 | -18.14716   9.404074   -1.93   0.054   -36.57881    .2844865
-----+-----
```

```
. margins, at(c1sess=(0(1)5) reas22=(-5 0 5)) vsquish // predictions per session
```

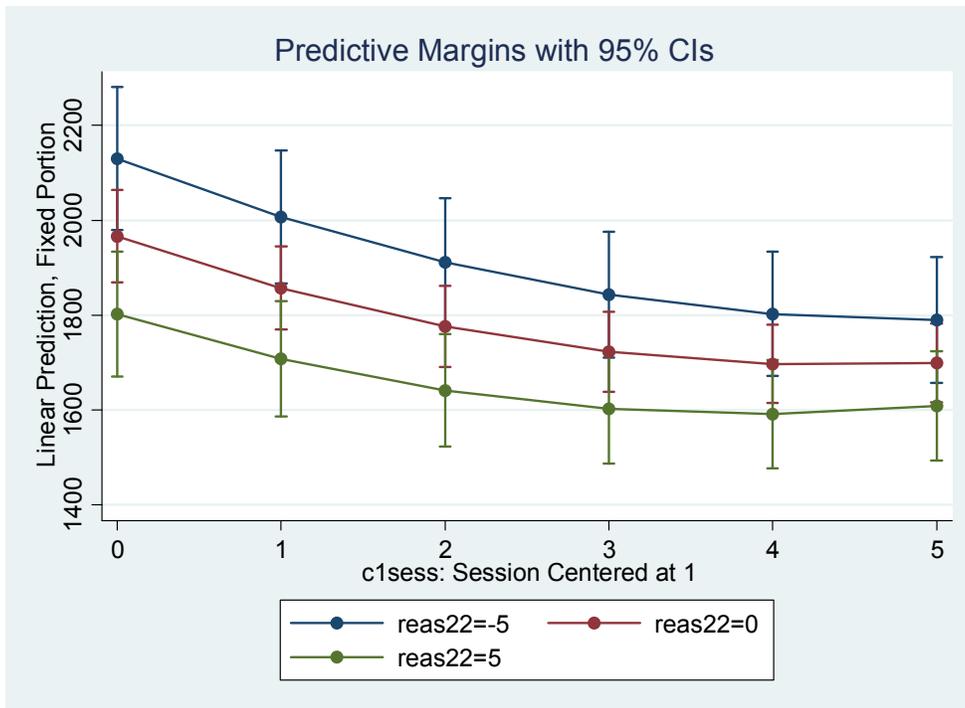
```
Predictive margins                                Number of obs    =          606
```

```
Expression   : Linear prediction, fixed portion, predict()
```

```
1._at       : c1sess      =      0
              reas22      =     -5
2._at       : c1sess      =      0
              reas22      =      0
3._at       : c1sess      =      0
              reas22      =      5
```

```
(output continues for all other sessions)
```

```
-----+-----
          |          Delta-method
          |          Margin  Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
   _at |
  1 | 2130.467   76.79056   27.74   0.000   1979.96    2280.974
  2 | 1966.327   49.71952   39.55   0.000   1868.878   2063.775
  3 | 1802.186   67.29827   26.78   0.000   1670.284   1934.089
  4 | 2006.92    71.31601   28.14   0.000   1867.143   2146.697
  5 | 1857.461   44.5668    41.68   0.000   1770.112   1944.81
  6 | 1708.002   62.03248   27.53   0.000   1586.42    1829.583
  7 | 1911.105   69.07653   27.67   0.000   1775.717   2046.492
  8 | 1776.326   43.59481   40.75   0.000   1690.882   1861.771
  9 | 1641.548   60.20767   27.26   0.000   1523.543   1759.553
 10 | 1843.021   67.77101   27.19   0.000   1710.192   1975.849
 11 | 1722.923   43.0615    40.01   0.000   1638.524   1807.322
 12 | 1602.825   59.15403   27.10   0.000   1486.886   1718.765
 13 | 1802.668   66.80286   26.98   0.000   1671.736   1933.599
 14 | 1697.251   41.95444   40.45   0.000   1615.022   1779.48
 15 | 1591.834   58.16663   27.37   0.000   1477.83    1705.839
 16 | 1790.046   67.57664   26.49   0.000   1657.598   1922.494
 17 | 1699.31    42.47414   40.01   0.000   1616.062   1782.558
 18 | 1608.574   58.8501    27.33   0.000   1493.23    1723.918
-----+-----
```



```

. marginsplot, name(predicted_reas, replace) // plot reas predictions
Variables that uniquely identify margins: c1sess reas22

. predict predreas2 // save fixed-effect predicted outcomes
(option xb assumed)

. corr nm3rt predreas2 // get total r to make r2
(obs=606)
    
```

	nm3rt	predre~2
nm3rt	1.0000	
predreas2	0.4001	1.0000

R = .4001, so R² for time+age+reas = .1601
 So ~0.1% of the variance accounted for previously was due to reason*quad

Model 6a. +Education Group on Intercept, Linear, and Quadratic Time Slopes

Level 1: $y_{ti} = \beta_{0i} + \beta_{1i} (\text{Session}_{ti} - 1) + \beta_{2i} (\text{Session}_{ti} - 1)^2 + e_{ti}$

Level 2:

Intercept: $\beta_{0i} = \gamma_{00} + \gamma_{01} (\text{Age}_i - 80) + \gamma_{02} (\text{Reason}_i - 22) + \gamma_{03} (\text{Highvs.LowEd}_i) + \gamma_{04} (\text{Highvs.MedEd}_i) + U_{0i}$

Linear: $\beta_{1i} = \gamma_{10} + \gamma_{11} (\text{Age}_i - 80) + \gamma_{12} (\text{Reason}_i - 22) + \gamma_{13} (\text{Highvs.LowEd}_i) + \gamma_{14} (\text{Highvs.MedEd}_i) + U_{1i}$

Quadratic: $\beta_{2i} = \gamma_{20} + \gamma_{21} (\text{Age}_i - 80) + \gamma_{23} (\text{Highvs.LowEd}_i) + \gamma_{24} (\text{Highvs.MedEd}_i) + U_{2i}$

Additional model-implied group differences:

Medium vs. Low education intercept = $(\gamma_{00} + \gamma_{04}) - (\gamma_{00} + \gamma_{03}) = \gamma_{04} - \gamma_{03}$

Medium vs. Low education linear session = $(\gamma_{10} + \gamma_{14}) - (\gamma_{10} + \gamma_{13}) = \gamma_{14} - \gamma_{13}$

Medium vs. Low education quadratic session = $(\gamma_{20} + \gamma_{24}) - (\gamma_{20} + \gamma_{23}) = \gamma_{24} - \gamma_{23}$

```

TITLE1 "SAS Model 6a: +Education Group on Intercept, Linear, and Quadratic";
PROC MIXED DATA=work.example23 NOCLPRINT NOITPRINT COVTEST NAMELEN=100 IC METHOD=ML;
  CLASS ID session educgrp;
  MODEL nm3rt = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
              reas22 c1sess*reas22 educgrp c1sess*educgrp c1sess*c1sess*educgrp
              / SOLUTION DDFM=Satterthwaite OUTPM=work.EducPred;
  RANDOM INTERCEPT c1sess c1sess*c1sess / G GCORR TYPE=UN SUBJECT=ID;
  REPEATED session / TYPE=VC SUBJECT=ID;
  * Estimating group means at first and last sessions
  LSMEANS educgrp / AT (c1sess) = (0) DIFF=ALL;
  LSMEANS educgrp / AT (c1sess) = (5) DIFF=ALL;
  * Contrasts between groups on intercept, linear, and quadratic slopes
  ESTIMATE "L vs. H Educ for Intercept Main Effect" educgrp -1 0 1 ;
  ESTIMATE "M vs. H Educ for Intercept Main Effect" educgrp 0 -1 1 ;
  ESTIMATE "L vs. M Educ for Intercept Main Effect" educgrp -1 1 0 ;
  ESTIMATE "L vs. H Educ for Linear Session" c1sess*educgrp -1 0 1 ;
  ESTIMATE "M vs. H Educ for Linear Session" c1sess*educgrp 0 -1 1 ;
  ESTIMATE "L vs. M Educ for Linear Session" c1sess*educgrp -1 1 0 ;
  ESTIMATE "L vs. H Educ for Quadratic Session" c1sess*c1sess*educgrp -1 0 1 ;
  ESTIMATE "M vs. H Educ for Quadratic Session" c1sess*c1sess*educgrp 0 -1 1 ;
  ESTIMATE "L vs. M Educ for Quadratic Session" c1sess*c1sess*educgrp -1 1 0 ;
RUN; PROC CORR NOSIMPLE DATA=work.EducPred; VAR nm3rt pred; RUN;

```

Think of the -1 as the "0" and the "1" as the "1" in a dummy code.

```

TITLE "SPSS Model 6a: +Education as Predictor of Intercept, Linear, and Quadratic".
MIXED nm3rt BY ID session educgrp WITH c1sess age80 reas22
  /METHOD = ML
  /PRINT = SOLUTION TESTCOV G R
  /FIXED = c1sess c1sess*c1sess age80 c1sess*age80 c1sess*c1sess*age80
           reas22 c1sess*reas22 educgrp c1sess*educgrp c1sess*c1sess*educgrp
  /RANDOM = INTERCEPT c1sess c1sess*c1sess | SUBJECT(ID) COVTYPE(UN)
  /REPEATED = session | SUBJECT(ID) COVTYPE(ID)
  /SAVE = FIXPRED (prededuc)
  /EMMEANS = TABLES(educgrp) WITH (c1sess=0) COMPARE(educgrp)
  /EMMEANS = TABLES(educgrp) WITH (c1sess=5) COMPARE(educgrp)
  /TEST = "L vs. H Educ for for Main Effect" educgrp -1 0 1
  /TEST = "M vs. H Educ for for Main Effect" educgrp 0 -1 1
  /TEST = "L vs. M Educ for for Main Effect" educgrp -1 1 0
  /TEST = "L vs. H Educ for for Linear Session" c1sess*educgrp -1 0 1
  /TEST = "M vs. H Educ for for Linear Session" c1sess*educgrp 0 -1 1
  /TEST = "L vs. M Educ for for Linear Session" c1sess*educgrp -1 1 0
  /TEST = "L vs. H Educ for for Quadratic Session" c1sess*c1sess*educgrp -1 0 1
  /TEST = "M vs. H Educ for for Quadratic Session" c1sess*c1sess*educgrp 0 -1 1
  /TEST = "L vs. M Educ for for Quadratic Session" c1sess*c1sess*educgrp -1 1 0.
CORRELATIONS nm3rt prededuc.

```

```

* STATA Model 6a: +Education Group on Intercept, Linear, and Quadratic
xtmixed nm3rt c.c1sess c.c1sess#c.c1sess ///
  c.age80 c.age80#c.c1sess c.age80#c.c1sess#c.c1sess ///
  c.reas22 c.reas22#c.c1sess ///
  b(last).educgrp ib(last).educgrp#c.c1sess ///
  ib(last).educgrp#c.c1sess#c.c1sess, || id: c1sess c1sess2, ///
  variance ml covariance(un) residuals(independent,t(session)),
  estat ic, n(101),
  estat recovariance, level(id),
  estimates store Educ,
  lrtest Educ Reas2,
  * Estimating group means at first and last sessions
  margins ib(last).educgrp, at(c.c1sess=(0 5))
  * Contrasts between groups on intercept, linear, and quadratic slopes
  test 1.educgrp=3.educgrp // Low vs. High: Intercept
  test 2.educgrp=3.educgrp // Med vs. High: Intercept
  test 1.educgrp=2.educgrp // Low vs. Med: Intercept
  test 1.educgrp#c.c1sess=3.educgrp#c.c1sess // Low vs. High: Linear
  test 2.educgrp#c.c1sess=3.educgrp#c.c1sess // Med vs. High: Linear
  test 1.educgrp#c.c1sess=2.educgrp#c.c1sess // Low vs. Med: Linear
  test 1.educgrp#c.c1sess#c.c1sess=3.educgrp#c.c1sess#c.c1sess // Low vs. High: Quad
  test 2.educgrp#c.c1sess#c.c1sess=3.educgrp#c.c1sess#c.c1sess // Med vs. High: Quad
  test 1.educgrp#c.c1sess#c.c1sess=2.educgrp#c.c1sess#c.c1sess // Low vs. Med: Quad

```

```

contrast educgrp, // omnibus group diff on intercept
contrast educgrp#c.c1sess, // omnibus group diff on linear
contrast educgrp#c.c1sess#c.c1sess, // omnibus group diff on quadratic
margins, at(c.c1sess=(0(1)5) educgrp=(1 2 3)) vsquish // predictions per session
marginsplot, name(predicted_educ, replace) // plot educ predictions
predict prededuc // save fixed-effect predicted outcomes
corr nm3rt prededuc // get total r to make r2
    
```

STATA output:

```

Mixed-effects ML regression      Number of obs   =      606
Group variable: id              Number of groups =      101
                                Obs per group: min =       6
                                avg =      6.0
                                max =       6
                                Wald chi2(13)   =     106.94
                                Prob > chi2     =      0.0000
Log likelihood = -4147.6829
    
```

	nm3rt	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	c1sess	-106.4987	40.28349	-2.64	0.008	-185.4529	-27.54452
	c.c1sess#c.c1sess	12.47966	6.848972	1.82	0.068	-.9440805	25.9034
	age80	20.28963	8.560341	2.37	0.018	3.511673	37.06759
	c.age80#c.c1sess	-4.575964	3.267261	-1.40	0.161	-10.97968	1.827749
	c.age80#c.c1sess#c.c1sess	.6176862	.5534216	1.12	0.264	-.4670003	1.702373
	reas22	-36.62127	10.76417	-3.40	0.001	-57.71865	-15.52389
	c.reas22#c.c1sess	2.978327	1.280262	2.33	0.020	.4690609	5.487594
	educgrp						
	1	-51.37682	151.0698	-0.34	0.734	-347.4683	244.7146
	2	37.64254	120.8739	0.31	0.755	-199.2659	274.5509
	educgrp#c.c1sess						
	1	-70.24589	59.07811	-1.19	0.234	-186.0368	45.54507
	2	-4.357662	48.13262	-0.09	0.928	-98.69587	89.98055
	educgrp#c.c1sess#c.c1sess						
	1	11.06526	10.03239	1.10	0.270	-8.597857	30.72837
	2	-1.464111	8.188464	-0.18	0.858	-17.51321	14.58498
	_cons	1961.886	101.7896	19.27	0.000	1762.382	2161.39

Random-effects Parameters	Estimate	Std. Err.	[95% Conf. Interval]		
id: Unstructured					
var(c1sess)	24143.68	5618.86	15300.54	38097.82	→ linear var down by 2.95%
var(c1sess2)	582.0323	164.4561	334.5305	1012.648	→ quad var down by 4.01%
var(_cons)	228602.6	34699.74	169776.2	307812.1	→ intercept var down by 0.04%
cov(c1sess,c1sess2)	-3636.004	939.9484	-5478.269	-1793.739	
cov(c1sess,_cons)	-33285.68	10917.47	-54683.52	-11887.83	
cov(c1sess2,_cons)	4127.595	1789.937	619.382	7635.808	
var(Residual)	20298.12	1649.105	17310.14	23801.87	→ residual var not reduced

```

LR test vs. linear regression:      chi2(6) =    826.31   Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.
    
```

```

.      estat ic, n(101),
    
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	101	.	-4147.683	21	8337.366	8392.283

Note: N=101 used in calculating BIC

Is the education model (6a) better than the revised reasoning model (5b)?
 No, $-2\Delta LL = 4.7$ on $df=6$, $p = .583$

```
. estimates store Educ,
. lrtest Educ Reas2,
```

```
Likelihood-ratio test                               LR chi2(6) =      4.70
(Assumption: Reas2 nested in Educ)                 Prob > chi2 =    0.5831
```

```
. * Estimating group means at first and last sessions
. margins ib(last).educgrp, at(c1sess=(0))
Predictive margins                                Number of obs   =      606
Expression   : Linear prediction, fixed portion, predict()
at           : c1sess = 0
```

```
-----+-----
```

		Delta-method				
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
educgrp						
1	1884.284	110.9362	16.99	0.000	1666.853	2101.715
2	1973.304	66.56898	29.64	0.000	1842.831	2103.776
3	1935.661	101.2482	19.12	0.000	1737.218	2134.104

```
-----+-----
```

```
. margins ib(last).educgrp, at(c1sess=(5))
Predictive margins                                Number of obs   =      606
Expression   : Linear prediction, fixed portion, predict()
at           : c1sess = 5
```

```
-----+-----
```

		Delta-method				
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
educgrp						
1	1599.713	94.54367	16.92	0.000	1414.41	1785.015
2	1704.939	56.6088	30.12	0.000	1593.988	1815.89
3	1725.687	86.06347	20.05	0.000	1557.006	1894.369

```
-----+-----
```

```
. * Contrasts between groups on intercept, linear, and quadratic slopes
. test 1.educgrp=3.educgrp // Low vs. High: Intercept
( 1) [nm3rt]1.educgrp - [nm3rt]3b.educgrp = 0
     chi2( 1) = 0.12
     Prob > chi2 = 0.7338

. test 2.educgrp=3.educgrp // Med vs. High: Intercept
( 1) [nm3rt]2.educgrp - [nm3rt]3b.educgrp = 0
     chi2( 1) = 0.10
     Prob > chi2 = 0.7555

. test 1.educgrp=2.educgrp // Low vs. Med: Intercept
( 1) [nm3rt]1.educgrp - [nm3rt]2.educgrp = 0
     chi2( 1) = 0.46
     Prob > chi2 = 0.4960

. test 1.educgrp#c.c1sess=3.educgrp#c.c1sess // Low vs. High: Linear
( 1) [nm3rt]1.educgrp#c.c1sess - [nm3rt]3b.educgrp#co.c1sess = 0
     chi2( 1) = 1.41
     Prob > chi2 = 0.2344

. test 2.educgrp#c.c1sess=3.educgrp#c.c1sess // Med vs. High: Linear
( 1) [nm3rt]2.educgrp#c.c1sess - [nm3rt]3b.educgrp#co.c1sess = 0
     chi2( 1) = 0.01
     Prob > chi2 = 0.9279

. test 1.educgrp#c.c1sess=2.educgrp#c.c1sess // Low vs. Med: Linear
```

```
( 1) [nm3rt]1.educgrp#c.c1sess - [nm3rt]2.educgrp#c.c1sess = 0
      chi2( 1) = 1.69
      Prob > chi2 = 0.1939

.      test 1.educgrp#c.c1sess#c.c1sess=3.educgrp#c.c1sess#c.c1sess // Low vs. High: Quad
( 1) [nm3rt]1.educgrp#c.c1sess#c.c1sess - [nm3rt]3b.educgrp#co.c1sess#co.c1sess = 0
      chi2( 1) = 1.22
      Prob > chi2 = 0.2700

.      test 2.educgrp#c.c1sess#c.c1sess=3.educgrp#c.c1sess#c.c1sess // Med vs. High: Quad
( 1) [nm3rt]2.educgrp#c.c1sess#c.c1sess - [nm3rt]3b.educgrp#co.c1sess#co.c1sess = 0
      chi2( 1) = 0.03
      Prob > chi2 = 0.8581

.      test 1.educgrp#c.c1sess#c.c1sess=2.educgrp#c.c1sess#c.c1sess // Low vs. Med: Quad
( 1) [nm3rt]1.educgrp#c.c1sess#c.c1sess - [nm3rt]2.educgrp#c.c1sess#c.c1sess = 0
      chi2( 1) = 2.12
      Prob > chi2 = 0.1454

.      contrast educgrp, // omnibus group diff on intercept
```

Contrasts of marginal linear predictions
 Margins : asbalanced

	df	chi2	P>chi2
nm3rt			
educgrp	2	0.48	0.7869

```
.      contrast educgrp#c.c1sess, // omnibus group diff on linear
Contrasts of marginal linear predictions
Margins : asbalanced
```

	df	chi2	P>chi2
nm3rt			
educgrp#c.c1sess	2	1.92	0.3827

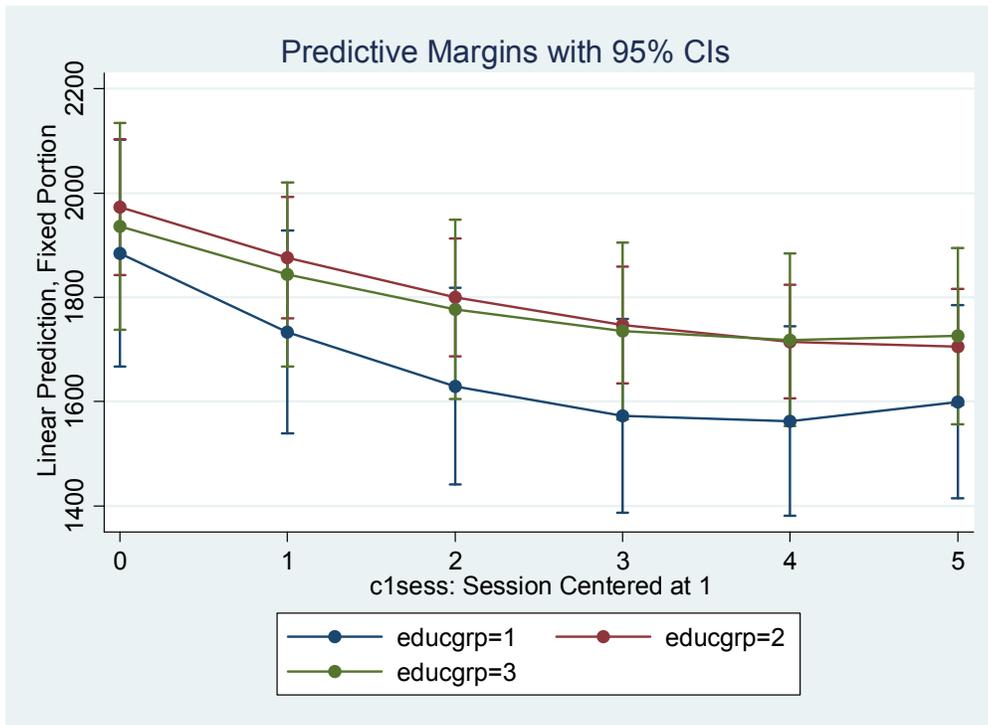
```
.      contrast educgrp#c.c1sess#c.c1sess, // omnibus group diff on quadratic
Contrasts of marginal linear predictions
Margins : asbalanced
```

	df	chi2	P>chi2
nm3rt			
educgrp#c.c1sess#c.c1sess	2	2.18	0.3358

```
.      margins, at(c.c1sess=(0(1)5) educgrp=(1 2 3)) vsquish // predictions per session
Predictive margins      Number of obs = 606
Expression : Linear prediction, fixed portion, predict()
1._at      : c1sess      = 0
            : educgrp    = 1
2._at      : c1sess      = 0
            : educgrp    = 2
3._at      : c1sess      = 0
            : educgrp    = 3
(output continues for all other sessions)
```

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]	
_at						
1	1884.284	110.9362	16.99	0.000	1666.853	2101.715
2	1973.304	66.56898	29.64	0.000	1842.831	2103.776
3	1935.661	101.2482	19.12	0.000	1737.218	2134.104
4	1733.602	98.99478	17.51	0.000	1539.576	1927.628
5	1875.98	59.24278	31.67	0.000	1759.867	1992.094
6	1844.159	90.05891	20.48	0.000	1667.647	2020.672
7	1629.804	96.28618	16.93	0.000	1441.086	1818.521
8	1800.482	57.64972	31.23	0.000	1687.491	1913.474
9	1777.411	87.64532	20.28	0.000	1605.63	1949.193
10	1572.889	94.93902	16.57	0.000	1386.812	1758.967
11	1746.809	56.86718	30.72	0.000	1635.352	1858.267
12	1735.417	86.46259	20.07	0.000	1565.953	1904.88
13	1562.859	92.79589	16.84	0.000	1380.983	1744.736
14	1714.961	55.5437	30.88	0.000	1606.098	1823.825
15	1718.175	84.43878	20.35	0.000	1552.678	1883.672
16	1599.713	94.54367	16.92	0.000	1414.41	1785.015
17	1704.939	56.6088	30.12	0.000	1593.988	1815.89
18	1725.687	86.06347	20.05	0.000	1557.006	1894.369

```
. marginsplot, name(predicted_educ, replace) // plot educ predictions
Variables that uniquely identify margins: c1sess educgrp
```



```
. predict prededuc // save fixed-effect predicted outcomes
(option xb assumed)
```

```
. corr nm3rt prededuc // get total r to make r2
(obs=606)
```

	nm3rt	prededuc
nm3rt	1.0000	
prededuc	0.4151	1.0000

R = .4151, so R² for time+age+reas+educ = .172

The fixed effects of time, age, and reasoning before accounted for ~16.0% of the variance in RT, so there is a net increase of 1.2% due to education (which is not significant).

Simple Processing Speed – Example Conditional Models of Change Results

The extent to which individual differences in response time (RT) over six sessions for a simple processing speed test (number match three) could be predicted from baseline age, abstract reasoning, and education level was examined in a series of multilevel models (i.e., general linear mixed models) in which the six practice sessions were nested within each participant. Given the interest in comparing models differing in fixed effects, maximum likelihood (ML) was used in estimating and reporting all model parameters. The significance of new fixed effects were evaluated with individual Wald tests (i.e., of estimate / SE) as well as with likelihood ratio tests (i.e., $-2\Delta LL$), with degrees of freedom equal to the number of new fixed effects. Session (i.e., the index of time) was centered at the first occasion, age was centered at 80 years, abstract reasoning was centered at 22 (near the mean of the scale), and graduate-level education was the reference group for education level (with separate contrasts for high school or less and for bachelor's level education).

The best-fitting unconditional growth model specified quadratic decline across the six sessions (i.e., a decelerating negative function) with significant individual differences in the intercept, linear, and quadratic effects. Accordingly, effect size was evaluated via pseudo- R^2 values for the proportion reduction in each random effect variance, as well as with total R^2 , the squared correlation between the actual outcome values and the outcomes predicted by the model fixed effects. In the unconditional growth model, the fixed effects for linear and quadratic change across sessions accounted for approximately 4% of the total variation in RT.

Next, age was added as a predictor of the intercept, linear slope, and quadratic slope. The age model fit significantly better than the unconditional model as indicated by a significant likelihood ratio test, $-2\Delta LL(3) = 11.6$, $p = .009$; the AIC was lower, although the BIC was not. However, only the fixed effect of age on the intercept was significant, indicating that for every additional year of age above 80, RT at the first session was predicted to be significantly higher by 29.05 ($p = .001$). In terms of pseudo- R^2 , age accounted for 11.29% of the random intercept variance, 4.50% of the random linear slope variance, and 2.64% of the random quadratic slope variance. As expected given that baseline age is a time-invariant predictor, the residual variance was not reduced. The total cumulative R^2 from session and age was $R^2 = .11$, approximately a 7% increase due to age (which was significant, as indicated by the likelihood ratio test). Although the interactions of age with the linear and quadratic slopes were not significant, they were retained in the model to fully control for age effects before examining the effects of other predictors.

Abstract reasoning was then added as a predictor of the intercept, linear slope, and quadratic slope. The abstract reasoning model fit significantly better than the age model, $-2\Delta LL(3) = 12.5$, $p = .006$; the AIC was lower, although the BIC was not. However, only the fixed effect of reasoning on the intercept was significant. The nonsignificant effect of reasoning on the quadratic slope was then removed, revealing a significant effect of reasoning on both the intercept and linear slope, such that for every unit higher reasoning above 22, RT at the first session was expected to be lower by 32.82 and the linear rate of improvement in RT (as evaluated at the first session given the quadratic slope) was expected to be less negative by 2.94 (i.e., faster initial RT with less improvement in persons with greater reasoning). These two effects still resulted in a significant improvement in model fit over the age model, $-2\Delta LL(2) = 10.1$, $p = .006$, with a lower AIC and BIC. Reasoning accounted for 5.68% of the random intercept variance but had no measurable reduction of the random linear and quadratic slope variances. The total cumulative R^2 from session, age, and reasoning was $R^2 = .16$, approximately a 5% increase due to reasoning (which was significant, as indicated by the likelihood ratio test).

Finally, education level (high school or less, bachelor's level, or graduate level) was then added as a predictor of the intercept, linear slope, and quadratic slope. The education model did not fit significantly better than the reasoning model, $-2\Delta LL(6) = 4.7$, $p = .583$, with a higher AIC and BIC. None of the omnibus main effects of group on the intercept, linear, or quadratic slopes were significant, $\chi^2(2) < 1.92$, p 's $> .05$, and none of the pairwise group comparisons were significant as well. Education accounted for 0.04% of the random intercept variance, 2.95% of the random linear slope variance, and 4.01% of the random quadratic slope variance. The total cumulative R^2 from session, age, reasoning, and education was $R^2 = .17$, approximately a 1% increase due to education (which was not significant, as indicated by the likelihood ratio test).

(From here one might remove nonsignificant model effects and/or add other effects as needed to fully answer all research questions...)